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Infinite divisibility
and the waiting-time paradox

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INFINITE DIVISIBILITY AND THE WAITING-TIME PARADOX

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Abstract. It is shown that the interarrival time $Z$ covering the point zero in a stationary renewal process generated by $X$ has the form $Z \overset{d}{=} X + Y$ with $Y$ nonnegative and independent of $X$, if and only if $X$ is infinitely divisible. In the special case that $X$ has a compound-exponential distribution there is a similar decomposition of the stationary waiting time. These results shed some new light on the waiting-time paradox.

1. Introduction and summary

The waiting-time paradox consists of the fact that a passenger arriving at a bus stop will probably have to wait considerably longer than about half the interarrival time, $X$ say, of two buses. The paradox is resolved by the observation that a passenger is more likely to arrive in a long interval than in a short one. Alternatively, the length $Z$ of the interval covering the arrival time of the passenger tends to be longer than $X$, i.e., we have

\[(1.1) \quad Z \geq X\]

in distribution. On an appropriate sample space (1.1) can be written as

\[(1.2) \quad Z \overset{d}{=} X + Y,\]

where $Y$ is nonnegative and, in general, $X$ and $Y$ are dependent.

In this note we consider the case where $X$ and $Y$ in (1.2) are independent. It turns out that this happens if and only if $X$ is infinitely divisible, and this result sheds some new light on the waiting-time paradox and on the behaviour of nonnegative processes with stationary, independent increments [sii-processes]. In the special case that $X$ is compound-exponential [see Definition 2.5], also the (stationary) waiting-time $W$ of a bus passenger admits of a decomposition similar to (1.2):

\[(1.3) \quad W \overset{d}{=} X + A,\]

where $A$ is nonnegative and independent of $X$.

In Section 2 we collect some results about renewal theory and infinite divisibility. Section 3 contains a characterization of infinite divisibility for nonnegative random variables by means of (1.2), and some of its consequences. In Section 4 equation (1.3) is considered for non-lattice distributions. In Section 5 we briefly present the solutions of (1.2) and (1.3) for distributions on the non-negative integers. We shall mostly use the renewal (life-time) terminology rather than the waiting-time terminology.
2. Preliminaries

We need some information on renewal theory and on infinite divisibility for non-negative random variables. Some of the results hold only for non-lattice distributions, and have analogues for lattice distributions. We shall only consider lattice random variables with values in $Z_+ := \mathbb{N} \cup \{0\}$. Sometimes it will be essential that these random variables have positive probability at 0.

2.1. Renewal theory

We consider a renewal process generated by a sequence $(X_n)_{n \in \mathbb{N}}$ of independent, non-negative random variables distributed as $X$ with distribution function $F$ and expectation $\mathbb{E}X = \mu \in (0, \infty)$; these conditions will be assumed to hold throughout unless otherwise stated. We write $S_n = \sum_{k=1}^{n} X_k$, and define the number $N_t$ of renewals in $(0, t]$, the age $V_t$ of the unit in service at time $t$, the remaining lifetime $W_t$ of this unit (the waiting time for the next bus), and the total life-time $Z_t$ as follows:

$$
N_t := \# \{ n \in \mathbb{N} : S_n \leq t \},
$$

$$
V_t := t - S_{N_t}, \quad W_t := S_{N_t+1} - t,
$$

$$
Z_t = X_{N_t+1} = V_t + W_t.
$$

The following result is well known; see Feller (1971).

**Lemma 2.1.** If $F$ is non-lattice, then

$$
V_t \overset{d}{\rightarrow} V, \quad W_t \overset{d}{\rightarrow} W \overset{d}{=} V, \quad Z_t \overset{d}{\rightarrow} Z \quad [t \to \infty],
$$

where $Z$ and $W$ have distribution functions given by

$$
F_Z(z) = \frac{1}{\mu} \int_{[0,z]} x \, dF(x) \quad [z \geq 0],
$$

$$
F_W(w) = \frac{1}{\mu} \int_{[0]}^w \{1 - F(x)\} \, dx \quad [w \geq 0].
$$

Alternatively, $V$, $W$ and $Z$ can be regarded as quantities in a stationary renewal process (started at $-\infty$); then $Z$ is the life time straddling 0, say, and $V$ and $W$ are the lengths of the parts into which $Z$ is divided by 0. The random variables $V$, $W$ and $Z$ will be used in this sense, sometimes without comment. They satisfy

$$
Z = V + W \quad \text{and} \quad V \overset{d}{=} W \overset{d}{=} U \, Z,
$$

where $U$ is uniformly distributed on $(0,1)$ and independent of $Z$; see Winter (1989).
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There is an analogue to Lemma 2.1 for lattice distributions. Now $X$ is distributed on $\mathbb{Z}_+$, and the quantities $N_n, V_n, W_n$ and $Z_n$ are defined for $n \in \mathbb{N}$ in the same way as $N_t, V_t, W_t$ and $Z_t$. We then have the following well-known result; see Feller (1968). We use the notations $V, W$ and $Z$ also for the limits of $V_n, W_n$ and $Z_n$.

**Lemma 2.2.** If $X$ has an aperiodic distribution $(p_k)_{k \in \mathbb{Z}_+}$ on $\mathbb{Z}_+$, then

$$V_n \xrightarrow{d} V, \quad W_n \xrightarrow{d} W \xrightarrow{d} V + 1, \quad Z_n \xrightarrow{d} Z \quad [n \to \infty],$$

where $Z$ and $W$ have distributions given by

\begin{align*}
(2.4) \quad \mathbb{P}(Z = k) &= \frac{1}{\mu} k p_k \quad [k \in \mathbb{N}], \\
(2.5) \quad \mathbb{P}(W = k) &= \frac{1}{\mu} \sum_{j=k}^{\infty} p_j \quad [k \in \mathbb{N}].
\end{align*}

**Remark.** Comparison with (2.4) learns that (2.1) also holds for distributions on $\mathbb{Z}_+$. Since the span of the lattice is of no real importance, (2.1) holds for all distributions on $\mathbb{R}_+$ with a finite moment. Since $W$ in (2.2) is absolutely continuous, the non-lattice condition there is essential. □

### 2.2. Infinite divisibility

We recall the definition of infinite divisibility: a random variable $X$ (or its distribution function $F$) is said to be **infinitely divisible** if for any $n \in \mathbb{N}$ iid random variables $X^{(n)}_1, \ldots, X^{(n)}_n$ exist such that

\begin{align*}
(2.6) \quad X \equiv X^{(n)}_1 + \cdots + X^{(n)}_n.
\end{align*}

In what follows the Laplace-Stieltjes transform [LSt] of a function $H$ will be denoted by $\tilde{H}$. We take the following result from Feller (1971).

**Lemma 2.3.** A distribution function $F$ on $\mathbb{R}_+$ is infinitely divisible iff its LSt satisfies

\begin{align*}
(2.7) \quad \log \tilde{F}(s) &= \int_0^\infty (e^{-sx} - 1) x^{-1} dK(x),
\end{align*}

where $K$ is a nondecreasing function on $\mathbb{R}_+$ with (necessarily) $\int_1^\infty x^{-1} dK(x) < \infty$. $K$ will be called the **canonical function** of $F$, or of the corresponding random variable. We note that the random variables $X^{(n)}$ in (2.6) are infinitely divisible with distribution function $F^{*(1/n)}$, the $n$-th convolution root of $F$, and with canonical function $K^{(n)} = (1/n) K$. Differentiation of (2.7) yields

\begin{align*}
(2.8) \quad -\frac{\tilde{F}'(s)}{\tilde{F}(s)} &= \tilde{K}(s),
\end{align*}

and inversion of this leads to the following equivalent relation [see Steutel (1970)]; let $s \downarrow 0$ for the final statement.
Corollary 2.4. A distribution function $F$ on $\mathbb{R}_+$ is infinitely divisible iff $F$ satisfies
\begin{equation}
\int_{(0,x]} y \, dF(y) = (F * K)(x) \quad [x > 0],
\end{equation}
where $K$ is nondecreasing on $\mathbb{R}_+$ and $*$ denotes convolution. The, possibly infinite, first moment $\mu$ of $F$ is given by
\begin{equation}
\mu = \int_{[0,\infty)} dK(x) = K(\infty).
\end{equation}

For Section 4 we need the following subclass of infinitely divisible distributions.

Definition 2.5. Let $S(\cdot)$ be a process with stationary, independent increments and let $T$ be exponentially distributed with $E T = 1$, independent of $S(\cdot)$. Then the random variable $X := S(T)$ is said to have a compound-exponential distribution.

The following lemma is immediate if we denote the (infinitely divisible) distribution function of $S(1)$ by $F_0$.

Lemma 2.6. A distribution function $F$ on $\mathbb{R}_+$ is compound-exponential iff $F$ has the form
\begin{equation}
\widehat{F}(s) = \frac{1}{1 - \log \widehat{F_0}(s)},
\end{equation}
where $F_0$ is an infinitely divisible distribution function on $\mathbb{R}_+$. The distribution functions $F$ and $F_0$ have the same first moment.

We conclude this section with analogues (and also special cases) of Corollary 2.4 and Lemma 2.6 for infinitely divisible distributions on $\mathbb{Z}_+$. Here we use probability generating functions [pgf's] rather than LSt's. We stress that the factors $X_j^{(n)}$ in (2.3) are $\mathbb{Z}_+$-valued iff $P(X = 0) > 0$; for details we refer to Steutel (1970).

Lemma 2.7. A distribution $(p_k)_{k \in \mathbb{Z}_+}$ on $\mathbb{Z}_+$ with $p_0 > 0$ is infinitely divisible iff the quantities $r_k$ with $k \in \mathbb{Z}_+$ defined by
\begin{equation}
(k + 1) p_{k+1} = \sum_{j=0}^{k} p_j r_{k-j} \quad [k \in \mathbb{Z}_+],
\end{equation}
are nonnegative, i.e., iff the coefficients of $R(z) := P'(z)/P(z)$ are nonnegative.

Lemma 2.8. A distribution on $\mathbb{Z}_+$ with pgf $P$ is compound-exponential iff $P$ has the form
\begin{equation}
P(z) = \frac{1}{1 - \log Q(z)} = \frac{1}{1 - \alpha(G(z) - 1)},
\end{equation}
where $Q$ is the pgf of an infinitely divisible distribution on $\mathbb{Z}_+$ with $Q'/Q = \alpha G'$, $\alpha > 0$ and $G$ a pgf.
3. An extreme case of the waiting-time paradox

We return to the renewal process generated by a nonnegative random variable \( X \) as described in Section 2.1. It is well known [see e.g. Ross (1970)] that for the total life-time \( Z \) with distribution function given by (2.1) we have

\[
Z \overset{d}{=} X,
\]

i.e., \( P(Z > x) \geq P(X > x) \) for \( x \in \mathbb{R}_+ \), or on a suitable sample space,

\[
Z \overset{d}{=} X + Y,
\]

where \( Y \) is nonnegative and, in general, not independent of \( X \). Here we are interested in the situation where \( X \) and \( Y \) are independent. The following theorem can also be read as a characterization theorem for infinitely divisible random variables on \( \mathbb{R}_+ \) with a finite first moment [compare Remark following Lemma 2.2].

**Theorem 3.1.** Let \( Z \) be the life time covering the point 0 in a stationary renewal process generated by \( X \). Then \( Z \) can be written as in (3.1) with \( X \) and \( Y \) independent iff \( X \) is infinitely divisible. The distribution function of \( Y \) is given by \( F_Y = (1/\mu)K \), where \( K \) is the canonical function of \( X \).

**Proof.** Let \( Z \) satisfy (3.1) with \( X \) and \( Y \) independent. Then by (2.1) we have

\[
F_Z(z) = \frac{1}{\mu} \int_{[0,z]} x \, dF(x) = (F * F_Y)(z) \quad [z \geq 0].
\]

Multiplying by \( \mu \) we see that \( F \) satisfies the functional equation (2.9) with \( K = \mu F_Y \), hence by Corollary 2.4 \( F \) is infinitely divisible. Conversely, if \( F \) is infinitely divisible, then by (2.10) the functional equation (2.9) can be written in the form (3.2) for some distribution function \( F_Y \), which means that in (3.1) \( X \) and \( Y \) can be taken independent. \( \square \)

**Remark.** By a suitable choice of \( K \), for \( F_Y = (1/\mu)K \) any distribution function is possible. If \( F(0) > 0 \), then \( X \) is compound-Poisson [cf. van Harn (1978)], i.e.,

\[
X \overset{d}{=} \tilde{X}_1 + \cdots + \tilde{X}_N,
\]

where \( \tilde{X}_1, \tilde{X}_2, \ldots \) are independent and distributed as \( \tilde{X} \) with distribution function \( G \), say, and \( N \) is Poisson distributed, independent of the \( \tilde{X}_j \)'s. Now \( F_Y \) takes the form

\[
F_Y(y) = \frac{1}{\mathbb{E}X} \int_{(0,y]} u \, dG(u),
\]

i.e., \( Y \) has the same distribution as the total life-time \( \tilde{Z} \) in the renewal process generated by \( \tilde{X} \). This can, roughly, be explained as follows. Since \( X \) is a Poisson sum
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of \(\tilde{X}_j\)'s, the (positive) \(X\)-interval covering a fixed point can be regarded as the sum of the \(\tilde{X}\)-interval covering this point and a (Poisson) number of independent \(\tilde{X}_j\)'s, i.e., as \(X + Z = X + Y\).

The question remaining is: what is the meaning of the random variable \(Y\) in (3.2). An answer can be obtained by considering the renewal process generated by \(X^{(n)}\) with distribution function \(F^{*(1/n)}\). Denoting the corresponding \(Z\)-random variable by \(Z^{(n)}\), and using the fact that \(X^{(n)}\) is infinitely divisible with canonical function \(K^{(n)} = (1/n)K\), by Theorem 3.1 we have

\[
Z^{(n)} = X^{(n)} + Y^{(n)},
\]

where the distribution function of \(Y^{(n)}\) is given by

\[
F_{Y^{(n)}} = \frac{1}{\mu} F_{X^{(n)}} K^{(n)} = \frac{n}{\mu} K^{(n)} = \frac{1}{\mu} K = F_Y,
\]

independent of \(n\). More generally, considering the sii-process \(X(\cdot)\) with \(X(1) = X\) and \(X(1/n) = X^{(n)}\), we obtain the following result.

**Theorem 3.2.** Let \(X\) be a nonnegative infinitely divisible random variable with finite first moment \(\mu\) and with canonical function \(K\). Let \(X(\cdot)\) denote the sii-process with \(X(1) = X\), and for \(t > 0\) let \(Z(t)\) be the \(Z\)-random variable in a stationary renewal process generated by \(X(t)\). Then

\[
(3.3)\quad Z(t) \overset{d}{=} X(t) + Y \quad [t > 0],
\]

where \(X(t)\) and \(Y\) are independent and \(F_Y = (1/\mu)K\), independent of \(t\).

Equation (3.3) tells us that, if we approach the point 0 with small steps, i.e., if \(t\) is small, then this point is passed with a non-small step of length at least \(Y\), in distribution. Also, though \(X(t) \overset{d}{\rightarrow} 0\), we have \(Z(t) \overset{d}{\rightarrow} Y \neq 0\) as \(t \downarrow 0\): we get an extreme case of the waiting-time paradox. For large \(t\), on the other hand, the term \(X(t)\) will dominate in the right-hand side of (3.3): the paradox, almost, disappears. Combining (2.3) with (3.3) yields the following result for the remaining life-time \(W\).

**Corollary 3.3.** Let \(X(\cdot)\) be as in Theorem 3.2, and for \(t > 0\) let \(W(t)\) be the \(W\)-random variable in a stationary renewal process generated by \(X(t)\). Then

\[
(3.4)\quad W(t) \overset{d}{=} UX(t) + UY \quad [t > 0],
\]

where \(U, X(t)\) and \(Y\) are independent, \(U\) is uniform on \((0,1)\), and \(Y\) is as in Theorem 3.2. In particular,

\[
(3.5)\quad W(t) \overset{d}{\rightarrow} UY \quad [t \downarrow 0].
\]
4. Compound-exponential life times

Since the compound-exponential distributions [cf. Definition 2.5] are a special subclass of the infinitely divisible distributions, one may expect the random variable $Y$ in (3.1) to have special properties, if $X$ is taken compound-exponential. Indeed we have the following result.

**Theorem 4.1.** If $X$ is a nonnegative, compound-exponential random variable with a finite first moment, then the life time $Z$ covering 0 in a stationary renewal process generated by $X$ satisfies

$$Z \overset{d}{=} X + X' + Y_0,$$

where $X' \overset{d}{=} X$, $Y_0$ is nonnegative, and $X$, $X'$ and $Y_0$ are independent. Equivalently, the random variable $Y$ in (3.1) satisfies $Y \overset{d}{=} X' + Y_0$ with $X'$ and $Y_0$ independent and independent of $X$.

**Proof.** By Lemma 2.6 the LSt $\hat{F}$ of $X$ has the form $\hat{F}(s) = 1/(1 - \log \hat{F}_0(s))$, where $F_0$ is an infinitely divisible distribution function on $\mathbb{R}_+$. From this equation and (2.8) it follows that the canonical function $K$ of $F$ is related to the canonical function $K_0$ of $F_0$ by $K = F * K_0$:

$$\hat{K}(s) = -\frac{d}{ds} \log \hat{F}(s) = \hat{F}(s) \{ -\frac{d}{ds} \log \hat{F}_0(s) \} = \hat{F}(s) \hat{K}_0(s).$$

Dividing by $\mu$, the first moment of both $F$ and $F_0$, and letting $Y_0$ be a random variable with distribution function $(1/\mu) K_0$, we see that $Y$ in (3.1) satisfies

$$F_Y = F * F_{Y_0},$$

so $Y$ can be obtained as $Y \overset{d}{=} X' + Y_0$ with $X'$ and $Y_0$ independent.

**Remark.** As in Theorem 3.2 one can consider the renewal process generated by $X(t)$. When $X = X(1)$ is compound-exponential, then so is $X(t)$ for $0 < t < 1$. Repeating the calculations for Theorem 4.1 with $X$ replaced by $X(t)$, we obtain

$$Z(t) \overset{d}{=} X(t) + X'(t) + Y_0(t) \quad [0 < t \leq 1],$$

where $Y_0(t) \overset{d}{=} X(1-t) + Y_0$ with $Y_0$ as in (4.1) independent of $X(1-t)$. So, we recover (3.3) with $Y \overset{d}{=} X + Y_0$ and $Y_0$ as in (4.1).

Choosing a uniformly distributed random variable $U$ independent of $X$, $X'$ and $Y_0$, we can rewrite (4.1) as

$$Z \overset{d}{=} (X + UY_0) + (X' + UY_0),$$
where $\bar{U} = 1 - U \overset{d}{=} U$. Since we also have $Z = V + W$ with $V \overset{d}{=} W$ [cf. Lemma 2.1 and (2.3)], this suggests that in this case

\[(4.3) \quad W \overset{d}{=} X + A,\]

where $X$ and $A$ are independent with $A$ of the form $A = UY_0$. This is the content of the next theorem. Since apart from some mass at zero the random variable $A$, and hence $W$, has an absolutely continuous distribution, we consider the non-lattice case here; cf. Remark following Lemma 2.2.

**Theorem 4.2.** If $X$ is compound-exponential and non-lattice with finite first moment $\mu$, then the stationary remaining life-time $W$ satisfies (4.3) with $X$ and $A$ independent, and $A$ of the form $A \overset{d}{=} UY_0$, where $Y_0$ is nonnegative, $U$ is uniform on $(0,1)$ and independent of $Y_0$.

**Proof.** Use Lemmas 2.1, 2.3 and 2.6. Since $X$ has LSt $\hat{F} = 1/(1 - \log \hat{F}_0)$ with $F_0$ infinitely divisible, the LSt of $W$ can be written as

\[
\hat{F}_W(s) = \frac{1 - \hat{F}(s)}{\mu s} = \hat{F}(s) - \frac{\log \hat{F}_0(s)}{\mu s} = \hat{F}(s) \int_0^\infty \frac{1 - e^{-sx}}{sx} dF_0(x),
\]

where, as in the proof of Theorem 4.1, $F_0 = (1/\mu) K_0$ with $K_0$ the canonical function of $F_0$. The equality in the extreme members of this equation easily translates in (4.3) with $A$ of the form $A = UY_0$ as desired. \(\square\)

It is unclear whether (4.3) with $X$ and $A$ independent implies that $X$ has a compound-exponential distribution. The fact that $A$ is of the form $A \overset{d}{=} UY_0$ implies that $A$ has a density on $(0,\infty)$ which is nonincreasing.

**Remark.** Proceeding as in the Remark following Theorem 4.1, for $W(t)$ we obtain

\[(4.4) \quad W(t) \overset{d}{=} X(t) + UY_0(t) \quad [0 < t \leq 1],\]

which can be written as

\[(4.5) \quad W(t) \overset{d}{=} (1 - U)X(t) + U\{X(t) + Y_0(t)\},\]

where $X(t) + Y_0(t) \overset{d}{=} Y$, independent of $t$; compare Corollary 3.3. \(\square\)

5. Analogues for lattice distributions

In this section we briefly present the analogues of Theorems 3.1, 4.1 and 4.2 for $\mathbb{Z}_+$-valued random variables $X$. We consider the discrete-time renewal process generated by $X$ with $\mathbb{P}(X = k) = p_k$ for $k \in \mathbb{Z}_+$ and $\mathbb{E}X = \mu \in (0,\infty)$, and use the
notation as established in Lemma 2.2. We restrict attention to $X$ with $p_0 > 0$. This leads to the clearest analogues, and it is no real restriction: The results for $Z$ hold for arbitrary $X$; for the result on $W$ the restriction $P(X \leq 1) > 0$ is necessary, and the case $p_0 = 0, p_1 > 0$ can be reduced to the case $p_0 > 0$ by a shift. The proofs differ only slightly from those in Sections 3 and 4.

**Theorem 5.1.** Let $X$ have an aperiodic distribution $(p_k)_{k \in \mathbb{Z}^+}$ on $\mathbb{Z}^+$ with $p_0 > 0$. Then the random variable $Z$ with distribution given by (2.4) satisfies

$$Z \overset{d}{=} X + Y,$$

where $Y$ is a (necessarily $\mathbb{N}$-valued) random variable independent of $X$, iff $X$ is infinitely divisible.

**Proof.** By (2.4), saying that $Z$ satisfies (5.1) with $X$ and $Y$ independent, is equivalent to the following assertion:

$$P(Z - 1 = k) = \frac{1}{\mu} (k + 1)p_{k+1} = \sum_{j=0}^{k} p_j P(Y - 1 = k - j) \quad [k \in \mathbb{Z}^+] .$$

An appeal to Lemma 2.7 finishes the proof; note that necessarily $\sum_{k=0}^{\infty} r_k = \mu$. \[\square\]

From the proof it will be clear that $Y$ in (5.1) has distribution $(r_{k-1}/\mu)_{k \in \mathbb{N}}$. Of course, also Theorem 3.2 has a lattice analogue; we don’t spell it out.

The analogues of Theorems 4.1 and 4.2 are combined in the following theorem. Note that a compound-exponential distribution on $\mathbb{Z}^+$ necessarily has positive mass at 0.

**Theorem 5.2.** Let $X$ have an aperiodic, compound-exponential distribution $(p_k)$ on $\mathbb{Z}^+$. Then the random variable $Z$ with distribution given by (2.4) satisfies

$$Z \overset{d}{=} X + X' + Y_0,$$

where $X' \overset{d}{=} X, Y_0$ is $\mathbb{N}$-valued, and $X, X'$ and $Y_0$ are independent. For the random variable $W$ with distribution given by (2.5) one has

$$W \overset{d}{=} X + A,$$

where $X$ and $A$ are independent and $A$ has a nonincreasing distribution on $\mathbb{N}$.

**Proof.** By Lemma 2.8 the pgf $P$ of $X$ has the form $P = 1/(1 - \log Q)$ with $Q$ an infinitely divisible pgf; hence differentiation of $1/P$ leads to the equation $P'/P^2 = Q'/Q$. Since the pgf $P_{Z-1}$ of $Z - 1$ is given by $(1/\mu) P'$, it follows that

$$P_{Z-1}(z) = \{P(z)\}^2 (1/\mu) Q'(z)/Q(z).$$
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Now, \((1/\mu) Q'/Q\) can be read as the pgf of \(Y_0 - 1\) for some \(\mathbb{N}\)-valued random variable \(Y_0\); cf. the remark following Theorem 5.1. We conclude that \(Z\) satisfies (5.2). With regard to (5.3) it is sufficient to note that

\[
P_{W-1}(z) = \frac{1 - P(z)}{\mu(1 - z)} = P(z) \frac{-\log Q(z)}{\mu(1 - z)} = P(z) \frac{1 - G(z)}{\mu(\alpha)(1 - z)};
\]

cf. the second representation of (2.13) in Lemma 2.8.

6. An example

For \(X(\cdot)\) in Theorem 3.2 we take a Gamma process: \(X(t)\) has a gamma distribution with parameters \(t\) and \(\lambda\), so with density \(x \mapsto \lambda^t x^{t-1} e^{-\lambda x} / \Gamma(t)\). The canonical function of \(X(t)\) has density \(x \mapsto t e^{-\lambda x}\), and it follows that for the total life-time \(Z(t)\) in the stationary renewal process generated by \(X(t)\) \([t > 0, \text{fixed}]\) we have

\[
Z(t) \overset{d}{=} X(t) + Y,
\]

where \(X(t)\) and \(Y\) are independent and \(Y\) has an exponential distribution with parameter \(\lambda\), independent of \(t\).

For \(0 < t \leq 1\) the random variable \(X(t)\) is compound-exponential: its LST can be represented as \(1/(1 - \log \hat{F}_0)\), where \(F_0\) is infinitely divisible with canonical function \(K_0\) satisfying

\[
\hat{K}_0(s) = -\frac{\hat{F}'_0(s)}{\hat{F}_0(s)} = -\frac{d}{ds} \log \hat{F}_0(s) = \frac{t}{\lambda} \left( \frac{\lambda}{\lambda + s} \right)^{1-t}.
\]

From Theorem 4.1 it follows that for \(0 < t \leq 1\)

\[
Z(t) \overset{d}{=} X(t) + X'(t) + Y_0(t),
\]

where \(Y_0(t)\) has a gamma distribution with parameters \(1 - t\) and \(\lambda\). The special case \(t = 1\), i.e., \(X(1)\) exponential, gives \(Y_0(1) = 0\) with probability 1, as is well known. In general, \(Z(t)\) has a gamma distribution with shape parameter \(t + t + (1 - t) = t + 1\); for \(t = \frac{1}{2}\) the random variable \(Z(t)\) is the sum of three iid random variables.

By Theorem 4.2, for the remaining life-time \(W(t)\) in the renewal process generated by \(X(t)\) we find

\[
W(t) \overset{d}{=} X(t) + A(t),
\]

with

\[
\hat{F}_{A(t)}(s) = \frac{(1 + s/\lambda)^t - 1}{ts/\lambda}.
\]

For \(t = 1\) we obtain the well-known result that \(W(1) \overset{d}{=} X(1)\), i.e., \(\mathbb{P}(A(1) = 0) = 1\).
References


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