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Interpreting the copolymerization of styrene with maleic anhydride and with methyl methacrylate in terms of the bootstrap model

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(Received 14 January 1992; revised 18 May 1992)

The bootstrap model was proposed by Harwood to explain copolymer sequence distribution which is dependent only on copolymer composition in solvents which cause changes in the apparent reactivity ratios for a number of pairs of monomers. This work presents an explicit formulation of the equations describing copolymer composition and sequence distribution in terms of the bootstrap model for both the simple Mayo-Lewis model and the penultimate unit model. Inspection of these equations reveals the necessary and sufficient conditions for sequence distribution to be a function only of copolymer composition. The results of copolymerizations of styrene with methyl methacrylate and with maleic anhydride under various conditions are interpreted in terms of the bootstrap model.

(Keywords: copolymerization; bootstrap model; Mayo-Lewis model; penultimate unit model; apparent reactivity ratios; distribution coefficient)

INTRODUCTION
Harwood has presented a bootstrap model which deals with solvent effects in free radical copolymerization. In the original paper by Harwood four comonomer pairs show a pronounced solvent effect on determined reactivity ratios when copolymer composition versus monomer feed is studied. All four comonomer pairs consisted of a polar and a non-polar monomer. However, when the copolymer microstructure versus copolymer composition is examined, no such solvent effect is found. A copolymer having the same composition appears to have the same microstructure irrespective of the solvent employed during polymerization. Harwood states that, if the above is the case, the conditional probabilities governing the propagation reaction must be independent of the solvent. The differences observed in the copolymer composition versus monomer feed plots are due to a phenomenon which he calls the bootstrap effect. This effect means that a growing polymer radical influences its own environment. The comonomer ratio available for the growing chain can therefore differ from the global comonomer ratio.

Davis re-analysed the data of San Roman et al., and showed that for a comonomer pair with a relatively small difference in polarity (styrene/methyl methacrylate) the bootstrap effect can also be observed. The sequence distributions upon which Davis based his examination were measured by n.m.r. spectroscopy. The few publications dealing with the bootstrap model provide only qualitative descriptions of the effect. The aim of this paper is to gain insight into the quantitative aspects of the bootstrap model. Both for the Mayo-Lewis model and for the penultimate unit model the mathematical derivation of the relationship between microstructure and polymer composition will be presented. Furthermore, a method will be proposed for quantification of the distribution coefficient \( K \) which was introduced by Harwood in the presentation of his bootstrap model.

THEORY
For the Mayo-Lewis model the following set of equations is given to describe polymer composition and microstructure as a function of the relevant conditional probabilities:

\[
\begin{align*}
F_1 = \frac{p_{21}}{p_{12}} \\
F_{111} &= (1 - p_{12})^2 \\
F_{112} + F_{211} &= 2p_{12}(1 - p_{12}) \\
F_{212} &= p_{12}^2
\end{align*}
\]
Bootstrap model in styrene copolymerization. B. Klumperman and K. F. O'Driscoll

in which: \( F_i \) is the fraction of monomer \( i \) in the copolymer; 
\( F_{ij} \) is the fraction of monomer 1 centred triads with 
\( \Sigma F_{ij} = 1 \). The conditional probabilities \( p_{12} \) and \( p_{21} \) are 
defined as relative propagation rates. For example, \( p_{12} \) is 
the rate of addition of monomer 2 to a chain ending 
with monomer 1, where \( p_{11} + p_{12} = 1 \). In terms of 
the bootstrap model \( p_{12} \) and \( p_{21} \) can be written as:

\[
p_{12} = (1 + r_1 K q)^{-1}
\]

\[
p_{21} = (1 + r_2 K q)^{-1}
\]

where

\[
q = \frac{[M_0]}{[M_1]} \quad \text{and} \quad K = \frac{[M_1]/[M_2]}{[M_0]/[M_2]}
\]

and \( r_1, r_2 \) have their usual meanings as ratios of 
propagation rate constants and \( [M_0] \) are monomer 
concentrations. Superscript \(^0\) refers to global monomer 
concentration whereas no superscript refers to concen-
tration at the site of propagation.

In any copolymerization in a solvent the reactivity 
ratios determined are apparent reactivity ratios, \( r_i^a \),

\[
r_i^a = r_i K^a
\]

\[
r_2^a = r_2 K^a
\]

and \( K^a \) is the distribution coefficient as defined in 
equation (7) in the solvent.

It should be emphasized that independence of the 
conditional probabilities upon solvent most probably 
occur with independence of the monomer ratio at the 
site of propagation upon solvent. This means that in 
theory a 'real' reactivity ratio could be defined related to 
the monomer ratio at the site of propagation. This 
monomer ratio cannot be determined easily. It is, 
however, unnecessary to define real reactivity ratios, since 
the only objective is to relate the apparent reactivity ratios 
to a reference state. For the sake of quantification 
we can arbitrarily define such a reference state. For 
convenience we define the distribution coefficient to be 
unity (\( K^b = 1 \)) in bulk copolymerization, although the 
global monomer ratio in a bulk copolymerization may 
also differ from the monomer ratio at the site of 
propagation.

It follows that when reactivity ratios, \( r_i^b \), have been 
determined in bulk:

\[
r_1^b = K^b
\]

\[
r_2^b = K^b
\]

and that independent of the solvent employed:

\[
r_1^b r_2^b = r_1^a r_2^a
\]

By eliminating the monomer concentrations from the 
equations (1), (5) and (6) the following equation can be 
obtained:

\[
p_{12}^2 (r_1 r_2 - 1) F_1 + p_{12} + F_1 - 1 = 0
\]

with \( r_1 \neq 0 \). When \( r_1 = 0 \): \( p_{12}^b = 1 \) and equation (13) 
becomes invalid; although, when \( r_2 = 0 \), the equation is 
still valid.

Solving this equation yields the following expression 
for \( p_{12}^e \):

\[
p_{12}^e = \frac{1 - [4F_1^2(1 - r_1 r_2) - 4F_1(1 - r_1 r_2) + 1]^{1/2}}{2F_1(1 - r_1 r_2)}
\]

with \( r_1 r_2 \neq 1 \).

It becomes clear that if the bootstrap model applies 
the triad distribution versus copolymer composition is 
independent of the solvent. This can be seen easily from 
the dependence of the triad distribution on \( p_{12}^e \), which is 
only dependent on the product \( r_1 r_2 \). This product was 
shown to be independent of solvent. It is interesting to 
note that the square root function in equation (14) 
equals the quantity \( \kappa \) defined by Stockmayer when he 
mathematically derived the distribution of chain lengths 
and compositions in copolymers.

For the penultimate unit model a similar derivation 
can be used and the following set of equations applies:

\[
F_1 = \frac{1 + p_{112}}{p_{112}^e}
\]

\[
F_2 = \frac{1 + p_{211}}{p_{211}^e}
\]

\[
F_{111} = p_{112}(1 - p_{112})/p_{112}^e
\]

\[
F_{211} = 2p_{112}p_{211}/p_{112}^e + p_{211}^e
\]

\[
F_{212} = p_{112}(1 - p_{211})/p_{112}^e + p_{211}^e
\]

The conditional probabilities in this case can be written 
in a way similar to the Mayo–Lewis model, using the 
distribution coefficient \( K \):

\[
p_{112}^e = \frac{1}{1 + r_1 K q}
\]

\[
p_{211}^e = \frac{r_2 K q}{1 + r_2 K q}
\]

\[
p_{221}^e = \frac{1}{1 + r_2 K q}
\]

\[
p_{122}^e = \frac{r_1 K q}{1 + r_1 K q}
\]

where \( K \) and \( q \) are defined in the same way as in the 
derivation for the Mayo–Lewis model. The reactivity 
ratios \( r_{ij} \) are defined as:

\[
r_{ij} = k_{ij}/k_{jk}
\]

where subscript \( i, j \) and \( k \) may be 1 or 2 and subscript 
\( k \neq \text{subscript} \ j \). From equations (19)–(22) it can be seen 
that the apparent reactivity ratios in a solvent are defined 
as:

\[
r_{11}^b = r_{11} K^b
\]

\[
r_{12}^b = r_{12} K^b
\]
If we define the distribution coefficient $K^b$ to be unity for bulk copolymerization, analogous to the derivation using the Mayo–Lewis model it follows in this case:

$$\frac{r_{11}}{r_{12}} = K^b \quad (26)$$

$$\frac{r_{12}}{r_{21}} = K^b \quad (27)$$

The derivation of the equations describing the triad distribution as a function of polymer composition for the penultimate unit model appears to be complicated if no constraints are applied. However, for styrene maleic anhydride (STY/MAnh) copolymers the assumption is allowed that no homopropagation of monomer 2 (MAnh) occurs. Thus, equation (15) simplifies to

$$\frac{F_1}{F_2} = 1 + \frac{p_{211}}{p_{112}} \quad (28)$$

With this simplification $p_{112}$ and $p_{211}$ can be expressed as a function of copolymer composition

$$p_{211} = \frac{F_1}{F_1 - F_2} + \frac{4(F_1 - 1)(r_{11} - 1)}{2(F_{12} - 1)} \quad (29)$$

$$p_{112} = \frac{-F_1 + (1 - F_2)}{2(F_1 - 1)(r_{12} - 1)} \quad (30)$$

This implies that if no homopropagation of monomer 2 is present the triad distribution versus copolymer composition is only dependent on $r_{11}/r_{21}$. Knowing that the simple Mayo–Lewis model is a special case of the penultimate unit model, the dependence of the triad distribution versus copolymer composition on the product $r_1 r_2$ (or in this case $r_{12} r_{21}$) should be present too. Due to the symmetry in the copolymerization models the third factor in determining the triad distribution versus copolymer composition has to be $r_{12}/r_{21}$. Numerical inspection of these conditions using a computer model of the penultimate unit model confirms the three values to be the necessary and sufficient conditions for the triad distribution versus copolymer composition to be fixed. With due regard to the bootstrap model as put forward in equations (24) and (25) it is obvious that $r_{11}/r_{21}$, $r_{12}/r_{21}$ and $r_{12}/r_{12}$ are independent of the distribution coefficient $K$.

We presently are aware that the bootstrap model cannot only apply to the Mayo–Lewis model, but also to the penultimate unit model and presumably others. Once it is clear that a comonomer pair behaves according to either Mayo–Lewis or penultimate unit model it is possible to determine distribution coefficients in a reference state as outlined above. Equations (10) and (11) should be used to determine $K$ values for the Mayo–Lewis model and equations (26) and (27) for the penultimate unit model.

### Table 1

<table>
<thead>
<tr>
<th>Solvent</th>
<th>$r_1$</th>
<th>$r_5$</th>
<th>$r_{w5}$</th>
<th>$K$ (from $r_1$)</th>
<th>$K$ (from $r_5$)</th>
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<tr>
<td>Benzene</td>
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<td>0.51</td>
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<td>0.20</td>
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<tr>
<td>Benzonitrile</td>
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<td>0.47</td>
<td>0.31</td>
<td>1.34</td>
<td>1.44</td>
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<tr>
<td>Bulk</td>
<td>0.66</td>
<td>0.47</td>
<td>0.21</td>
<td>1.21</td>
<td>1.00</td>
</tr>
</tbody>
</table>

*Polymer composition determined by n.m.r.
**Polymer composition determined by carbon analysis

### Re-analysis of data on STY/MMA

STY/MMA is known to behave according to the Mayo–Lewis model with respect to copolymer composition and microstructure. Data by San Roman et al. is re-analysed using the error in variables method. The calculated values for the apparent reactivity ratios and their product is shown in Table 1.

The distribution coefficient can be calculated using equations (10) and (11) and the values for the reactivity ratios in bulk as reported by O'Driscoll et al. The distribution coefficient calculated from $r_1$ and $r_2$ are also given in Table 1.

The qualitative conclusion of Davis that the results of STY/MMA copolymerizations can be interpreted using the bootstrap model can now be quantified. It appears that the distribution coefficient $K$ increases when changing the solvent from benzene < bulk < chlorobenzene < benzonitrile. The physical background of this phenomenon has been reported earlier. Without involving the bootstrap model Russo correlated the variation in reactivity ratio to the polarity of the solvent. Now, taking into account the bootstrap model, one may conclude that the difference in polarity between solvent and site of polymerization (polymer coil) is the determining factor with respect to the distribution coefficient.

### STY/MAnh copolymerizations

For the STY/MAnh comonomer pair three different copolymerization conditions were studied. There is a high tendency to form alternating copolymers and, therefore, composition drift can occur very readily. It is for this reason that copolymerizations at a higher temperature were carried out using continuous copolymerization in a stirred tank reactor. Steady state conditions were used to ensure a constant comonomer ratio and thus a uniform copolymer composition. A disadvantage of the continuous copolymerization route, at least for this comonomer pair, is that the monomer ratio in the reactor is difficult to determine. However, for experiments where copolymer microstructure is studied as a function of copolymer composition the method appears to be valuable. For the determination of the apparent reactivity ratios in the case of STY/MAnh precipitation polymerization in toluene, an approach similar to the one described by Tidwell and Mortimer was used. This work will be described elsewhere.

### EXPERIMENTAL

#### Materials

For the batch experiments styrene was distilled at reduced pressure under nitrogen, to remove the inhibitor.
All solvents employed were dried using the usual techniques and stored over molecular sieves. All other chemicals were high purity grades and were used without further purification. For the batch experiments the required amounts of monomers, solvent and initiator were accurately weighed into the glass reaction vessels, thoroughly freeze–thaw degassed, and closed under vacuum. The polymerizations were carried out at 60°C to conversions less than 3 wt%. As the polymerizations were carried out in toluene, the copolymers precipitated on formation. The polymers were purified twice by reprecipitation from THF solution into isopropanol. The purified polymers were dried in vacuo at 50°C for 16 h.

For the continuous copolymerizations the required amounts of monomers, solvent and initiator were weighed into a storage vessel and kept under nitrogen. The mixture was pumped into a stainless steel reactor which was maintained at 110°C. After three times the mean residence time a polymer sample was collected. The polymers were recovered from solution and purification was carried out twice by reprecipitation from THF solution into isopropanol. The purified copolymers were dried in vacuo at 50°C for 16 h.

**Characterization**

DEPT 13C n.m.r. spectra were recorded on a Varian XL300 spectrometer at ambient temperature. Solutions were approximately 10% w/v in acetone D_6 for samples with high MAnh content or in CDCl_3 for samples with lower MAnh content. The compositions of the copolymers were determined from the triad fractions determined by DEPT 13C n.m.r. and using i.r. spectroscopy, from the ratio between the relative extinctions _E_155 cm⁻¹/ _E_1490 cm⁻¹. The latter method was calibrated against non-aqueous titration. The DEPT 13C n.m.r. technique to determine styrene centred triad distributions has been described earlier

**RESULTS**

The experimental results are summarized in Table 2. The styrene centred triad distribution versus copolymer composition is shown in Figure 1, together with data from Hill et al., who studied the copolymerization of STY/MAnh at 60°C in bulk. The data all seem to fit on the same set of curves, thus indicating the occurrence of the bootstrap effect. From equations (9)-(11), (15) and (16) it appears that the copolymer triad distribution versus composition only depends on the quotient _r_11/_r_21. This means that as in the STY/MMA case, where only the product _r_1/r_2 could be determined, only this quotient can be determined. Carrying out a non-linear fitting procedure gives _r_11/_r_21 = 0.63. The triad distribution calculated from equations (16)-(18), (29) and (30) for this value is shown in Figure 1 as drawn curves.

**DISCUSSION**

The copolymerization of STY/MAnh has received considerable attention in the literature. Various models have been used to try to describe this copolymerization. The complex participation model and the penultimate

<table>
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<tr>
<th>Solvent</th>
<th>f(MAnh) feed</th>
<th>f(MAnh) n.m.r.</th>
<th>Composition i.r.</th>
<th>SSS</th>
<th>MSS + SSM</th>
<th>MSM</th>
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<td>-</td>
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</table>
unit model appear to be the best in fitting the experimental data. In this work only the penultimate unit model is used. Since Harwood suggested that the bootstrap effect may cause or influence previously reported penultimate effects, the triad distribution versus copolymer composition data were also tested against the classical Mayo–Lewis model. However, if $r_2 = 0$ (no homopropagation of MAnh), then the product $r_1 r_2 = 0$ and therefore the triad distribution versus copolymer composition is fixed. This model shows poor agreement with the experimental data (dashed curves in Figure 1). As shown, the experimental data can be described well using the penultimate unit model with $r_{11}/r_{21} = 0.63$.

Apparent reactivity ratios were estimated from unpublished data on continuous polymerization in butanone at 110°C. The calculated values were $r_{11} = 0.20$ and $r_{21} = 0.33$. Although these values, compared with values reported in literature, show that reactivity ratios vary considerably with temperature, no effect is observed with respect to the ratio $r_{11}/r_{21}$. This may originate from activation energy of the various propagation rate constants and/or from the temperature dependence of the distribution coefficient. While the terminal monomer unit of the propagating chain and the incoming monomer are the same, the activation energy of $k_{11}$ and $k_{22}$ will be approximately the same. There may however exist a difference between $k_{12}$ and $k_{11}$. Therefore, it is possible that $r_{11}$ and $r_{21}$ differ, but that the quotient remains constant. On the other hand it is known that there is only a slight influence of temperature on apparent reactivity ratios. If, in the extreme case, reactivity ratios may be regarded as independent of temperature, values for the distribution coefficient may be calculated for butanone at 110°C. It is found that $K = 11.3$ can be calculated from $r_{11}$ and $K = 11.5$ from $r_{21}$. Reference values for bulk copolymerization were taken from the work by Hill et al. ($r_{11} = 0.0177; r_{21} = 0.0288$).

From the copolymer composition versus monomer feed data on the polymerizations in toluene, apparent reactivity ratios were calculated using a non-linear least squares method. It was found that $r_{11} = 0.021$ and $r_{21} = 0.049$ or $r_{11}/r_{21} = 0.43$. If reactivity ratios are also calculated from the five experiments where monomer feed ratio, copolymer composition and triad distribution are known the values are $r_{11} = 0.022$ and $r_{21} = 0.029$ or $r_{11}/r_{21} = 0.76$. The data is scattered for polymerizations in toluene, shown by the error bars in Figure 1 for two of the samples. The polymer precipitates on formation in toluene. Therefore, the scatter may be an artifact caused by the heterogeneity of the system. An additional source of error may arise from the low MAnh content in the monomer feed of some of the experiments, which readily leads to composition drift. If, notwithstanding the uncertainty in the reactivity ratio values, $K$-values for toluene are calculated from the latter $r_{11}$ and $r_{21}$ (since these are based on the largest amount of information), the following values are estimated. Calculated from $r_{11}$ it is found that $K = 1.2$ and calculated from $r_{21}$ it is found that $K = 1.0$.

CONCLUSIONS

From the present work it can be concluded that quantification of the bootstrap effect is possible. It should be emphasized that the choice of bulk copolymerization as reference state does not necessarily mean that in this case global comonomer ratio equals monomer ratio at the site of propagation. The bootstrap effect can apply to the Mayo–Lewis model and the penultimate unit model. In either case the copolymer microstructure versus composition data does not provide information about the individual reactivity ratio values or about the distribution coefficient. Such data only provide information about certain reactivity ratio product or quotient values.

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